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AUTHOR -

Barton, Mark A.; Lord, Frederic M.

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An upper-asymptote parameter was added to the three-parameter logistic item response model. This four-parameter model was compared to the three-parameter model on four data sets. The fourth parameter increased the likelihood in only two of the four sets. Ability estimates for the students were generally unchanged by the introduction of the fourth parameter. (Author)

AN UPPER ASYMPTOTE FOR THE THREE-PARAMETER

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Mark A. Barton

; and '

Frederic M. Lord

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Frederic M. Lord, Principal Investigator



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An Upper Asymptote for the Three-Parameter Logistic Item-Response Model

Mark A. Barton
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An Upper Asymptote for the Three-Parameter Logistic Kem-Response Model*

Mark A. Barton

and

Frederic M. Lord

A two-parameter logistic item-response model may be expressed as

$$F(\theta) = (1 + e^{-1.7a}g^{(\theta-bg')})^{-1}$$

where $F(\theta)$ is the probability of a student with ability θ passing item g. The probability of passing ranges from zero to one as θ goes from $-\infty$ to ∞ .

On a multiple choice test, however, the probability of choosing the / correct answer does not approach zero for low-ability students. For this reason, a lower asymptote is introduced. The resulting three-parameter model is

$$G(\theta) = c + (1 - c)\dot{F}(\theta)$$

where the lower asymptote, sometimes referred to as a guessing parameter, may be set arbitrarily; or a common value may be estimated from the data; or individual values may be estimated from the data separately for each item.

An upper asymptote may likewise be introduced into the model:

$$P(\theta) = c + (\delta - c)F(\theta)$$

While $F(\theta)$ ranges from zero to one, $P(\theta)$ ranges from the lower asymptote, c , to the upper asymptote, δ .

^{*}This work was supported in part by contract NO0014-80-C-0402, project designation NR150-453 between the Office of Naval Research and Educational Testing Service. Reproduction in whole or in part is permitted for any purpose of the United States Government.

Even a high-ability student may make a clerical error in answering an easy item. The introduction of an upper asymptote with a value of slightly less than I should allow a high-ability student to miss an easy item without having his ability estimate drastically lowered.

This study of the upper-asymptote model was motivated largely by the concern that the three-parameter model might be severely penalizing high-ability students who make a clerical error on an easy item. The three-parameter model was compared to the upper-asymptote (four-parameter) model for four data sets: SAT Verbal, SAT Math, GRE Verbal, and AP Calculus AB.

The tests, all of which were developed by Educational Testing
Service, are described briefly:

Scholastic Aptitude Test (SAT), Verbal: The SAT is a 75-minute, 90-item verbal test designed for applicants to undergraduate study.

Scholastic Aptitude Test (SAT), Mach: The mathematical part of the SAT (from a different form of the test, with different students) was analyzed as a separate data set. This 90-minute test has 85 items, 20 of which are four choice, and 65 of which are five choice.

Graduate Record Examination (GRE) Aptitude Test, Verbal: This test of verbal reasoning and reading comprehension is designed for applicants to graduate study. All of the items on this 50-minute, 80-item test are five-choice items.

College Board Advanced Placement (AP) Examination--Mathematics:

Calculus AB? This is a test taken by high school students in order to receive college credit for their knowledge of elementary functions and first-semester calculus. The essay items were not included in this analysis. The objective portion lasts 90 minutes and has 45 item.

All of the items, are five choice.

For each of the four data sets, the following procedure was used:

- .1) Ability and item parameters were estimated under the three-parameter model using the program LOGIST (Wood & Lord, 1976; Wood, Wingersky, & Lord, 1976), with several thousand students. (These runs were already available.)
 - 2) In order to reduce costs, a random subsample of 1000 students was chosen.
 - 3) With the parameters (guessing parameters) held fixed at the previously estimated values (to reduce costs), the remaining item parameters (a 's and b 's)' and the abilities (θ 's) were reestimated for just the 1000 students with δ = 1.00.
 - 4) With the c parameters still held fixed, the a 's, b 's, and θ 's were estimated with upper asymptote values of $\delta = .99 \quad \text{and} \quad \delta = .98 \quad \text{, using a modified version of}$ LOGIST on the sample of 1000 students.

The effect of the upper asymptote was measured in two ways. First, likelihoods were computed for each value of the upper asymptote (δ = 1.00 , δ = .99 , δ = .98), and the asymptote value with the highest likelihood was considered the best fit. In two of the cases (GRE Verbal and AP Calculus AB) the highest likelihood was obtained with δ = 1.00 (the standard three-parameter model) and in the other two cases (SAT Math and SAT Verbal) the highest likelihood. was obtained with δ = .99 . In none of the four data sets did the highest and the lowest of the three log-likelihoods differ by more than 0.02% (more than .0002).

Second, ability estimates under the two asymptote values δ = 1.00 and δ = .99 were compared to see whether the change in upperasymptote value had a large effect on the ability estimate of any individual student. Out of the 4000 students examined in the four data sets, the greatest change was an increase of .64 under the .99 upper-asymptote model. The five students most affected are summarized in Table 1. Two of these were very high-ability students, and the other three were very low-ability students. The ability estimate changed by more than the standard error of estimate in only one of the five cases. In all five cases, the .99-upper-asymptote model favored the student. The largest drop in a student's ability estimate under the .99-upper-asymptote model was .36, for a student of very low ability on the GRE Verbal. All 1000 students who took the GRE are plotted

Table 1

. The Five Students with the Largest Absolute Changes

in Theta out of 4000 Examined

		9			Standard Error of Measurement			
$\frac{\theta}{\delta = 1.00}$	$\delta = .99$		$\frac{\theta}{\delta = .99^{-\theta}} \delta = 1.00$		of $\theta_{\delta}=1.00$	_	Test	
2•. 9	3.4		0.5		0.4	٠.	GRE Verbal	
. 2.8	3.3		0.5 . •	٠	0.7		. AP Calculus	ΑB
. \-3.9	- 3.3		0.6		. 1.1		SAT Verbal	
-5.0	-4.4		. 0.6 .		2.7		GRE Verbal.	
-5.9	~ -5.4		· _ · 0.5		, 4.8		GRE Verbal	
			.		- ,		•	•

Thetas have a mean of $\boldsymbol{0}$ and a standard deviation of $\boldsymbol{1}$.

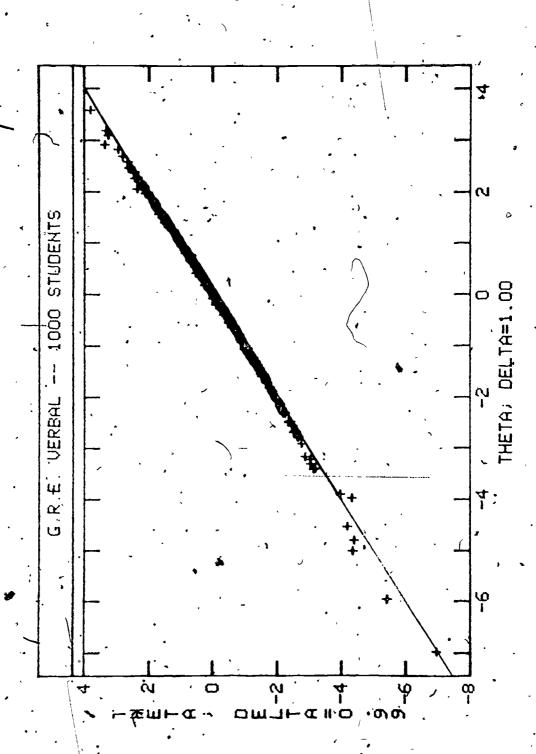
in Figure 1; in the plots for the other three data sets, points were even closer to the identity line.

One might expect more clerical errors to be made on a highly speeded test. This might result in the .99-upper-asymptote model fitting better than the 1.00-upper-asymptote model. The only highly speeded test of the four, the GRE Verbal, was fit best with the 1.00-upper-asymptote model, Nowever, discounting this speededness hypothesis.

In view of the failure of the four-parameter model either to consistently improve the likelihood or to significantly change any ability estimates there is no compelling reason to urge the use of this model. The extra computational time required for the more complex derivatives further argues against its use.

One positive result of this study is that it suggests that the three-parameter logistic model does a better job of fitting high-ability students than some researchers (including the authors) had expected. Had we examined the normal-ogive model instead, which is less forgiving of incorrect answers to easy items by high-ability students, we would possibly have found a stronger case for the adoption of an upper asymptote of .99 or .98.





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Appendix

Formulas for derivatives of the four-parameter model*:

n = number of items.

N = number of subjects.

 $\dot{D} = 1.7$.

 $\delta = \text{upper asymptote.}$

u = observed response (1 if right, 0 if wrong) for item i, subject j

 $L_{ij} = a_{i}(\theta_{j} - b_{i})$

 $P_{ij} = c_{i} + (\delta - c_{i})/(l_{*} + e^{-DL_{ij}})$

 $Q_{ij} = i - P_{ij}$

The log likelihood is given by

Theta derivatives:

$$\frac{\partial \ell}{\partial \theta_{j}} = \sum_{i=1}^{n} (u_{ij} - P_{ij}) \frac{(\sqrt{\partial \theta_{j}})}{P_{ij}Q_{ij}}$$

and

$$\frac{\frac{\partial^{2} \ell}{\partial \theta_{j}^{2}} = \sum_{i=1}^{n} \left[\frac{-\left(\frac{\partial P_{ij}}{\partial \theta_{j}^{i}}\right)^{2}}{P_{ij}Q_{ij}} + \left(u_{ij} - P_{ij}\right) \left[\frac{\left(\frac{\partial^{2} P_{ij}}{\partial \theta_{j}^{2}}\right)}{P_{ij}Q_{ij}} + \frac{\left(\frac{\partial^{2} P_{ij}}{\partial \theta_{j}^{2}}\right)^{2}}{P_{ij}Q_{ij}^{2}} + \frac{\left(\frac{\partial^{2} P$$

^{*}The authors are grateful to Kirsten Yocom for verifying all of the derivatives.

where

$$\frac{\partial P_{ij}}{\partial \theta_{j'}} = \frac{(\delta - c_{i})Da_{i}}{DL_{ij} + 2 + e}$$

and

$$\frac{\partial^{2} P_{ij}}{\partial \theta_{j}^{2}} = \frac{-(\delta - c_{i})D^{2}a_{i}^{2}(e^{DL_{ij}} - e^{-DL_{ij}})}{\frac{DL_{ij}}{(e^{DL_{ij}} + 2 + e^{-DL_{ij}})^{2}}}$$

- Item-parameter derivatives:

We will drop subscripts for brevity in the following formulas.

If χ represents a_i , b_i , or c_i , and ψ represents a_i , or c_i , then

$$\frac{\partial \mathcal{L}}{\partial \chi} = \sum_{j=1}^{N} (u - P) \frac{\left(\frac{\partial P}{\partial \chi}\right)}{PQ}$$

and

$$\frac{\partial^{2} \chi}{\partial \chi \partial \psi} = \sum_{j=1}^{N} \left[\frac{-(\frac{\partial P}{\partial \psi})(\frac{\partial P}{\partial \chi})}{PQ} + (u - P) \left[\frac{(\frac{\partial^{2} P}{\partial \chi \partial \psi})}{PQ} + (\frac{\partial P}{\partial \chi})(\frac{\partial P}{\partial \psi}) \frac{(P - Q)}{P^{2}Q^{2}} \right] \right]$$

where

$$\frac{\partial P}{\partial a} = \frac{(\delta - c)D(\theta' - b)}{e^{DL} + 2 + e^{-DL}}.$$

$$\frac{\partial P}{\partial b} = \frac{-(\delta - c)Da}{e^{DL} + 2 + e^{-DL}}$$

$$\frac{\partial P}{\partial c} = \frac{1}{1 + e^{DL}},$$

$$\frac{\partial^{2} P}{\partial a^{2}} = \frac{-(\delta - c)D^{2}(\theta - b)^{2}(e^{DL} - e^{-DL})}{(e^{DL} + 2 + e^{-DL})^{2}}.$$

$$\frac{\partial^2 P}{\partial b^2} = \frac{-(\delta - c)D^2 a^2 (e^{DL} - e^{-DL})}{(e^{DL} + 2 + e^{-DL})^2},$$

$$\frac{\partial^2 P}{\partial x^2} = 0$$

$$\frac{\partial^{2} P}{\partial a \partial b} = \frac{(\delta - c) D^{2} a (\theta - b) (e^{DL} - e^{-DL})}{(e^{DL} + 2 + e^{-DL})^{2}} - \frac{D(\delta - c)}{(e^{DL} + 2 + e^{-DL})}$$

$$\frac{\partial^2 P}{\partial a \partial c} = \frac{-D(\theta - b)}{e^{DL} + 2 + e^{-DL}}$$

and

$$\frac{\partial^2 P}{\partial h u c} = \frac{Da'}{e^{DL} + 2 + e^{-DL}}$$

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 Box 39
 FPO, NY 09510
- 1 Dr. Robert Brennan
 American College Testing Programs
 P.O. Box 168
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 - Dr. C. Victor Bunderson WICAT Inc. University Plaza, Suite 10 1160 South State Street Orem, UT 84057
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 Champaign, IL 61801
- 1 Dr. Charles Lewis
 Faculteit Sociale Wetenschappen
 Rijksuniversiteit Groningen
 Oude Boteringestraat 23
 9712GC Groningen
 NETHERLANDS
- Dr. Robert Lind
 College of Education
 University of Illinois
 Urbana, IL 61801
- 1 Dr. Frederic M. Lord Educational Testing Service Princeton, NJ 08541
- 1 Mr. Merl Malehorn
 Department of Navy
 Chief of Naval Operations
 OP-113
 Washington, DC 20350
- 1 Dr. Gary Marco Educational Testing Service Princeton, NJ 08541
- 1 Dr. Scott Maxwell
 Department of Psychology
 University of Houston
 Houston, TX 77004
- Dr. Samuel T. Mayo Loyola University of Chicago 820 North Michigan Avenue Chicago, IL 60611

- l Mr. Bill Nordbrock Instructional Program Development Building 90 NET-PDCD Great Lakes NTC, IL 60088
- Dr. Melvin R. Novick 356 Lindquist Center for Measurement University of Iowa Iowa City, IA 52242
- l Dr. Jesse Örlansky Institute for Defense Analyses 400 Army Navy Drive Arlington, VA 22202
- 1 Dr. James A. Paulson Portland State University P.O. Box 751 Portland, OR 97207
- 1 Mr. Luigi Petrul'lo 2431 North Edgewood Street Arlington, VA 22207
- l Dr. Diane M. Ramsey-Klee R-K Research and System Design 3947 Ridgemont Drive Malibu, CA 90265
- 1 Mr. Minrat M. L. Rauch
 P.II 4
 Bundesministerium der Verteidigung
 Postfach 1328
 D-53 Bonn 1
 GERMANY
- Dr. Mark D. Reckase
 Educational Psychology Department
 University of Missouri-Columbia
 4 Hill Hall
 Columbia, MO 65211
- 1 Dr. Andrew M. Rose American Institutes for Research 1055 Thomas Jefferson Street, NW . Washing on, DC 20007

- Dr. Leonard L. Rosenbaum, Chairman Department of Psychology Montgomery College Rockville, MD 20850
- Dr. Ernst Z. Rothkopf Bell Laboratories 600 Mountain Avenue Murray Hill, NJ 07974
- Dr. Lawrence Rudner
 403 Elm Avenue
 Takoma Park, MD 2001
- l Dr. J. Ryan
 Department of Education
 University of South Carolina
 Columbia, SC 29208
- Prof. Fumike Samejima
 Department of Psychology
 University of Tennessee
 Knoxville, TN 37916
- 1 Dr. Robert J. Seidel Instructional Technology Group HumRRO ** 300 North Washington Street Alexandria, VA 22314
- l Dr. Kazuo Shigemasu University of Tohoku Department of Educational Psychology Kawauchi, Sendai 980 JAPAN
- Dr. Edwin Shirkey Department of Psychology University of Central Florida Orlando, FL 32816
- 1 Dr. Robert Smith
 Department of Computer Science
 Rutgers University
 New Brunswick, NJ 08903

- 1 Dr. Richard Snow School of Education Stanford University Stanford, CA 94305
- 1 Dr. Robert Sternberg (Department of Psychology Yale University Box 11A, Yale Station New Haven, CT 06520
- 1 Dr. Patrick Suppes
 Institute for Mathematical Studies
 in the Social Sciences
 Stanford University.
 Stanford, CA 94305
- l Dr. Hariharan Swaminathan Laboratory of Psychometric and Evaluation Research School of Education University of Massacuusetts Amherst, MA 01003
- 1 Dr. Kikumi Tatsuoka
 Computer Bosed Education Research
 Labora Try
 252 Engineering Research Laboratory
 University of Illinois
 Urbana, IL 61801
- 1 Dr. David Thissen Department of Psychology University of Kansas Lawrence, KS 66044
- Dr. Robert Tsutakawa Department of Statistics University of Missouri Columbia, MO 65201
- 1 Dr. J. Uhlaner Perceptronics, Inc. 6271 Variel Avenue Woodland Hills, CA 91364

- 1 Dr. Howard Wainer Educational Testing Service Princeton, NJ 08541.
- Dr. Phyllis Weaver
 Graduate School of Education
 Harvard University
 200 Larsen Hall, Applan Way
 Cambridge, MA 02138
- Dr. David J. Weiss
 N660 Elliott Hall
 University.of Minnesota
 75 East River Road
 Minneapolis, MN 55455
- l Dr. Susan E. Whitely Psychology Department University:of Kansas Lawrence, KS 66044
- 1 Dr. Wolfgang Wildgrube Streitkraefteamt Box 20 50 03 D-5300 Bonn 2 WEST GERMANY